# LONG RANGE ORDER AND GIANT COMPONENTS OF QUANTUM RANDOM GRAPHS 

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#### Abstract

Mean field quantum random graphs give a natural generalization of classical Erdős-Rényi percolation model on complete graph $\mathrm{G}_{N}$ with $p=\beta / N$. Quantum case incorporates an additional parameter $\lambda \geqslant 0$, and the short-long range order transition should be studied in the $(\beta, \lambda)$-quarter plane. In this work we explicitly compute the corresponding critical curve $\gamma_{c}$, and derive results on two-point functions and sizes of connected components in both short and long range order regions. In this way the classical case corresponds to the limiting point $\left(\beta_{c}, 0\right)=(1,0)$ on $\gamma_{c}$.


## 1. Introduction and Results

1.1. Classical Erdős-Rényi random graphs. In the classical Erdős - Rényi model of random graphs each two vertices $i \neq j$ of the complete graph $\mathrm{G}_{N}=\{1, \ldots, N\}$ are connected with probability $p=\beta / N$ independently from all other edges. The phase transition [JLR] occurs at the critical value

$$
\begin{equation*}
\beta_{c}=1 \tag{1.1}
\end{equation*}
$$

Namely, for $\beta>\beta_{c}$ with probabilities of order $1-\mathrm{o}(1)$ there is a unique giant connected component of size $O(N)$, whereas for $\beta<\beta_{c}$ all the connected components of $\mathrm{G}_{N}$ have sizes of the order $O(\log N)$ or less.
1.2. Quantum version of Erdős-Rényi random graphs. Let us formulate now a quantum version of Erdős-Rényi random graphs. As we shall briefly explain in the end of the section, both the motivation and the choice of terminology comes from the stochastic geometric (Fortuin-Kasteleyn type) representation of quantum Curie-Weiss model in transverse magnetic field, which was originally developed in the general ferromagnetic context in [AKN].

There are two parameters $\beta \in(0, \infty]$ - the inverse temperature and $\lambda \in[0, \infty)$ - the strength of the transversal field. The case $\beta=\infty$ corresponds to the ground state, and the case $\lambda=0$ brings us back to the context of classical random graphs discussed above.

Given $\beta \in[0, \infty]$ let us use $\mathbb{S}_{\beta}$ to denote the circle of length $\beta$ under the convention $\mathbb{S}_{\infty}=\mathbb{R}$. The model is built on the space $\mathfrak{G}_{N}^{\beta}=\mathrm{G}_{N} \times \mathbb{S}_{\beta}$, that is to each site $i \in \mathrm{G}_{N}$ we attach a copy $\mathbb{S}_{\beta}^{i}$ of $\mathbb{S}_{\beta}$. With a slight abuse of notation we shall also write $\mathbb{S}_{\beta}^{i}=i \times \mathbb{S}_{\beta}$.

Our next step is to make finite $(\beta<\infty)$ or countable $(\beta=\infty)$ random number of holes in each $\mathbb{S}_{\beta}^{i}$ and draw finite $(\beta<\infty)$ or countable $(\beta=\infty)$ random number of

[^0]links between various points of $\mathfrak{G}_{N}$ with the same time coordinates, that is between points of the type $(i, t)$ and $(j, t)$ where $i \neq j$ and $t \in \mathbb{S}_{\beta}$. Both operations are going to be performed with the help of independent Poisson point processes over the time space $\mathbb{S}_{\beta}$ and, eventually, will lead to a splitting of $\mathfrak{G}_{N}^{\beta}$ into a finite $(\beta<\infty)$ or countable $(\beta=\infty)$ number of disjoint maximal connected components,
\[

$$
\begin{equation*}
\mathfrak{G}_{N}^{\beta} \backslash \mathcal{H}=\mathfrak{C}_{1} \vee \cdots \vee \mathfrak{C}_{n} \tag{1.2}
\end{equation*}
$$

\]

where $\mathcal{H}$ is the set of the holes. An example for $N=3$ is given on Figure 1.2. For each fixed $x \in \mathfrak{G}_{N}^{\beta}$ the probability $\mathbb{P}_{N}^{\beta, \lambda}(x \in \mathcal{H})=0$. Thus, for given $x \in \mathfrak{G}_{N}^{\beta}$ the notion $\mathfrak{C}(x)$ of the connected component containing $x$ in the decomposition (1.2) is well defined.


Figure 1. An example of the decomposition of $\mathfrak{G}_{3}^{\beta}$ after all the holes are punched and all the links are drawn: $\mathfrak{G}_{3}^{\beta} \backslash \mathcal{H}=\mathfrak{C}_{1} \vee \mathfrak{C}_{2}$, where $\mathfrak{C}_{1}=I_{1}^{1} \cup I_{1}^{2} \cup I_{3}^{1}$ and $\mathfrak{C}_{2}=I_{1}^{3} \cup I_{2}^{1} \cup I_{2}^{2} \cup I_{3}^{2}$

Processes of holes $\mathcal{H}_{i}$. For each $i \in G_{N}$ the process of holes $\mathcal{H}_{i}$ is the Poisson point process on $\mathbb{S}_{\beta}^{i}$ with intensity $\lambda . \mathcal{H}_{i}$-s are assumed to be independent for different $i$-s. For $\beta<\infty$ the punched circle $\mathbb{S}_{\beta}^{i} \backslash \mathcal{H}_{i}$ consists of $n$ disjoint connected intervals,

$$
\begin{equation*}
\mathbb{S}_{\beta}^{i} \backslash \mathcal{H}_{i}=I_{i}^{1} \cup I_{i}^{2} \cup \cdots \cup I_{i}^{n} \tag{1.3}
\end{equation*}
$$

Of course $n=1$ whenever the cardinality $\# \mathcal{H}_{i}=0,1$.

In the case $\beta=\infty$ the punched real line $\mathbb{S}_{\beta}^{i} \backslash \mathcal{H}_{i}$ is split into a countable disjoint union of connected intervals,

$$
\begin{equation*}
\mathbb{S}_{\infty}^{i} \backslash \mathcal{H}_{i}=\bigcup_{r=-\infty}^{\infty} I_{i}^{r} \tag{1.4}
\end{equation*}
$$

where we label $I_{i}^{0}$ the interval which contains $(i, 0)$.
In the sequel we shall use $\left|I_{i}^{k}\right|$ to denote the length of $I_{i}^{k}$, and we shall write $\mathcal{H}$ for the total collection of all the holes,

$$
\mathcal{H}=\cup_{i} \mathcal{H}_{i} \subset \mathfrak{G}_{N}^{\beta} .
$$

Processes of links $\mathcal{L}_{i j}$ and decomposition (1.2). With each (unordered) pair of vertices $i, j \in \mathrm{G}_{N}$ we associate a copy $\mathbb{S}_{\beta}^{i j}$ of $\mathbb{S}_{\beta}$ and a Poisson point process $\mathcal{L}_{i j}$ on $\mathbb{S}_{\beta}^{i j}$ with intensity $1 / N$. Processes $\mathcal{L}_{i j}=\mathcal{L}_{j i}$ are assumed to be independent for different $(i, j)$ and also independent of the processes of holes $\mathcal{H}_{i}$.

Two intervals $I_{i}^{k}$ and $I_{j}^{l}$ either in the decomposition (1.3) or accordingly, in the case $\beta=\infty$, in the decomposition (1.4) are said to be connected if there exists $t \in \mathbb{S}_{\beta}^{i j}$ such that $t \in \mathcal{L}_{i j}$, whereas $(i, t) \in I_{i}^{k}$ and $(j, t) \in I_{j}^{l}$. The decomposition (1.2) of $\mathfrak{G}_{N}^{\beta} \backslash \mathcal{H}$ into maximal connected components is, thereby, well defined.
Relation to the classical Erdős-Rényi random graph. If $\lambda=0$ then there are no holes and $\mathbb{S}_{\beta}^{i} \backslash \mathcal{H}_{i}$ always contains only one connected component, which of course equals to $\mathbb{S}_{\beta}^{i}$ itself. In the latter case, the probability $(\beta<\infty)$ that $\mathbb{S}_{\beta}^{i}$ and $\mathbb{S}_{\beta}^{j}$ are connected equals to $1-\mathrm{e}^{-\beta / N}$ and we are back to the original Erdős- Rényi setup.
1.3. Phase transition in the $(\beta, \lambda)$-plane. The critical curve $\gamma_{c}$ in the $(\beta, \lambda)$-coordinate quarter plane is implicitly given by (see Figure 1.3)

$$
\begin{equation*}
F(\beta, \lambda) \triangleq \frac{2}{\lambda}\left(1-\mathrm{e}^{-\lambda \beta}\right)-\beta \mathrm{e}^{-\lambda \beta}=1 \tag{1.5}
\end{equation*}
$$

It is easy to check that $\gamma_{c}$ is in fact a graph of a function $\lambda_{c}=\lambda_{c}(\beta)$ defined on $\beta \in[1, \infty)$. Consider the decomposition (Figure 1.3) of the off-critical region

$$
\mathbb{R}_{+}^{2} \backslash \gamma_{c}=A_{\mathrm{LRO}} \cup A_{\mathrm{SRO}}
$$

where

$$
\begin{equation*}
A_{\mathrm{LRO}}=\left\{(\beta, \lambda) \in \mathbb{R}_{+}^{2}: F(\beta, \lambda)>1\right\} \tag{1.6}
\end{equation*}
$$

LRO and SRO above stand for the long (respectively short ) range order.
Our main result states that for $(\beta, \lambda) \in A_{\text {LRO }}$ there is a long range order in the sense that the probability of two points $(i, j)$ and $(j, s)$ being connected does not vanish when the size of the system tends to infinity. Contrary to this such probability vanishes in the $N \rightarrow \infty$ limit whenever $(\beta, \lambda) \in A_{\text {SRO }}$. The survival of probabilities of connections is related to an emergence of an $O(N)$-giant connected component in the disjoint decomposition (1.2) in the $L R O$ regime; in particular for $(\beta, \lambda) \in A_{\text {SRO }}$ typical connected component of any point $(i, t) \in \mathfrak{G}_{N}^{\beta}$ is of order $O(\log N)$. This is a quantum version of Erdős- Rényi phase transition phenomenon and, since

$$
\lim _{\lambda \rightarrow 0} F(\beta, \lambda)=\beta
$$

the classical case is recovered in the limiting $\lambda=0$ case.
We proceed with several exact alternative statements of this result.


Figure 2. Decomposition of the $(\beta, \lambda)$ quarter plane into the short range and long range regions.
1.4. Long and short range order. In the sequel we shall use $\mathbb{P}_{N}^{\beta, \lambda}(\cdot)$ for the joint product measure of all the processes $\mathcal{H}_{i}$ of holes and all the processes $\mathcal{L}_{i j}$ of links as defined above.

Let us say that two points $(i, t),(j, s) \in \mathfrak{G}_{N}^{\beta}$ are connected if they belong to the same connected component in the decomposition (1.2). We shall denote the latter event as $\{(i, t) \longleftrightarrow(j, s)\}$.

Theorem A. If $(\beta, \lambda) \in A_{\mathrm{SRO}}$, then

$$
\begin{equation*}
\mathbb{P}_{N}^{\beta, \lambda}((i, t) \longleftrightarrow(j, s))=O\left(\frac{\log N}{N}\right) \tag{1.7}
\end{equation*}
$$

uniformly in $t, s \in \mathbb{S}_{\beta}$ and $i \neq j$.
On the other hand, if $\beta<\infty$ and $(\beta, \lambda) \in A_{\mathrm{LRO}}$, then there exists $p=p(\beta, \lambda) \in(0,1)$, such that

$$
\begin{equation*}
\mathbb{P}_{N}^{\beta, \lambda}((i, t) \longleftrightarrow(j, s))=p(\beta, \lambda)^{2}(1+\mathrm{o}(1)) \tag{1.8}
\end{equation*}
$$

also uniformly in $t, s \in \mathbb{S}_{\beta}$ and $i \neq j$.
1.5. Emergence of the giant component. Each connected cluster $\mathfrak{C}_{k}$ in the decomposition (1.2) consists of disjoint union of intervals

$$
\mathfrak{C}_{k}=\bigcup_{l} J_{k}^{l}
$$

where each $J_{k}^{l}$ coincides with some $I_{i}^{r}$ in one of the decompositions (1.3) ((1.4) in the $\beta=\infty$ case). Define,

$$
\left|\mathfrak{C}_{k}\right|=\sum_{l}\left|J_{k}^{l}\right| .
$$

For any fixed $(i, t) \in \mathfrak{G}_{N}^{\beta}$ the probability

$$
\mathbb{P}_{N}^{\beta, \lambda}((i, t) \in \mathcal{H})=0
$$

Thus, in general position, $(i, t) \in \mathfrak{G}_{N}^{\beta} \backslash \mathcal{H}$ and there exists $\mathfrak{C}_{k}(i, t)$ (which from now on we shall denote as $\mathfrak{C}((i, t))$ such that $(i, t) \in \mathfrak{C}_{k}$. Evidently the distribution of $|\mathfrak{C}(x)|$ does not depend on a particular $x=(i, t)$.

If $\beta<\infty$ we also define the maximal cluster size

$$
\begin{equation*}
\mathcal{M}=\max _{k}\left|\mathfrak{C}_{k}\right| \tag{1.9}
\end{equation*}
$$

and the next to maximal cluster size,

$$
\begin{equation*}
\mathcal{M}^{\text {next }}=\max \left\{\left|\mathfrak{C}_{k}\right|:\left|\mathfrak{C}_{k}\right| \neq \mathcal{M}\right\} . \tag{1.10}
\end{equation*}
$$

These definitions would clearly make little sense if $\beta=\infty$.
Theorem B. If $(\beta, \lambda) \in A_{\mathrm{SRO}}$, then for every $\kappa>0$ there exists $c=c(\beta, \lambda, \kappa)<\infty$, such that

$$
\begin{equation*}
\mathbb{P}_{N}^{\beta, \lambda}(|\mathfrak{C}(x)|>c \log N)=o\left(\frac{1}{N^{\kappa}}\right) \tag{1.11}
\end{equation*}
$$

Furthermore, if $\beta<\infty$, then

$$
\begin{equation*}
\mathbb{P}_{N}^{\beta, \lambda}(\mathcal{M}>c \log N)=o\left(\frac{1}{N^{\kappa-1}}\right) \tag{1.12}
\end{equation*}
$$

If, however, $\beta<\infty$ and $(\beta, \lambda) \in A_{\mathrm{LRO}}$ then there exists a sequence of positive numbers $\epsilon_{N}(\beta, \lambda) \rightarrow 0$ such that,

$$
\begin{equation*}
\mathbb{P}_{N}^{\beta, \lambda}\left(\left|\frac{|\mathfrak{C}(x)|}{N}-\rho\right|<\epsilon_{N}\right)=p(\beta, \lambda)(1-\mathrm{o}(1)) \tag{1.13}
\end{equation*}
$$

where $\rho=\rho(\beta, \lambda)=\beta p(\beta, \lambda)>0$, and $p(\beta, \lambda)$ is the same probability as in (1.8). Furthermore, in the $\beta<\infty$ case, there exists a constant $c=c(\beta, \lambda)<\infty$ such that

$$
\begin{equation*}
\mathbb{P}_{N}^{\beta, \lambda}\left(\mathcal{E}\left(\rho, \epsilon_{N}, c\right)\right)=1-\mathrm{o}(1) \tag{1.14}
\end{equation*}
$$

where the event $\mathcal{E}\left(\rho, \epsilon_{N}, c\right)$ is defined via

$$
\begin{equation*}
\mathcal{E}\left(\rho, \epsilon_{N}, c\right)=\left\{\left|\frac{\mathcal{M}}{N}-\rho\right|<\epsilon_{N}\right\} \cap\left\{\mathcal{M}^{\text {next }}<c \log N\right\} \tag{1.15}
\end{equation*}
$$

In this paper we shall prove the short range order parts of Theorem A and Theorem B for all $\beta \leqslant \infty$, whereas in the long range order case we shall concentrate only on the case of positive temperatures $\beta<\infty$. A treatment of the LRO properties of the ground state case $\beta=\infty$ requires an additional coupling with branching random walks in the sense of $[\mathrm{Bi}]$. Specifically, in the $\beta=\infty$ case in order to show that two large connected clusters intersect one should also control the time spread-off of each of these clusters. The corresponding results may be interesting in their own right and we relegate them to the forthcoming [IL].
1.6. Relation to Curie-Weiss model in transverse field. The Hamiltonian of the Curie-Weiss model in transversal field is given by

$$
\begin{equation*}
-\frac{1}{2 N} \sum_{i \neq j} \sigma_{i}^{\mathrm{z}} \sigma_{j}^{\mathrm{z}}-\lambda \sum_{i} \sigma_{i}^{\mathrm{x}} \tag{1.16}
\end{equation*}
$$

where $\sigma_{i}^{\mathrm{Z}}, \sigma_{i}^{\mathrm{X}}$ are Pauli spin $1 / 2$ matrices. As it has been discovered in [AKN] (in a general ferromagnetic context) path-integral type representation of the Curie-Weiss leads to the
following modification $\tilde{\mathbb{P}}_{N}^{\beta, \lambda}$ of our basic product measure of "holes" and "links", which is in fact the Fortuin-Kasteleyn representation of (1.16)

$$
\begin{equation*}
\tilde{\mathbb{P}}_{N}^{\beta, \lambda}(\mathrm{d} \mathcal{H}, \mathrm{~d} \mathcal{L})=\frac{1}{\mathcal{Z}_{N}(\beta, \lambda)} 2^{\# \mathrm{c}(\mathcal{H}, \mathcal{L})} \mathbb{P}_{N}^{\beta, \lambda}(\mathrm{d} \mathcal{H}, \mathrm{~d} \mathcal{L}) \tag{1.17}
\end{equation*}
$$

where $\#_{\mathrm{c}}(\mathcal{H}, \mathcal{L})$ is the number of maximal connected components in the decomposition (1.2). In particular, the two point function could be expressed in terms of this FK-measure as

$$
\left\langle\sigma_{i}^{Z} \sigma_{j}^{Z}\right\rangle_{N}^{\beta, \lambda}=\tilde{\mathbb{P}}_{N}^{\beta, \lambda}(i \longleftrightarrow j)
$$

A sample path large deviation analysis [IL] of (1.16) indicates that the long/short range order critical curve in the $(\beta, \lambda)$-plane is still given by (1.5). This, in view of the analysis of classical FK models on complete graphs [BGJ], is not very surprising, however so far we did not find a way it deduce it from purely stochastic geometric considerations, which would generalize recolouring techniques of the latter paper.
1.7. FKG properties of $\mathbb{P}_{N}^{\beta, \lambda}(\cdot)$. Many of our arguments rely on the following FKG (Fortuin-Kasteleyn-Ginibre) property of $\mathbb{P}_{N}^{\beta, \lambda}(\cdot)$ :

Let us define the partial order of the probability space, $\Omega$, in which $(\mathcal{H}, \mathcal{L})$ takes values in the following way:

$$
\left(\mathcal{H}^{\prime}, \mathcal{L}^{\prime}\right) \gg(\mathcal{H}, \mathcal{L}) \Leftrightarrow \mathcal{H}^{\prime} \subseteq \mathcal{H} \text { and } \mathcal{L}^{\prime} \supseteq \mathcal{L}
$$

In other words, in the decomposition of $\mathfrak{G}_{N}^{\beta}$ generated by $\left(\mathcal{H}^{\prime}, \mathcal{L}^{\prime}\right)$ there are less holes and more links than in the decomposition corresponding to $(\mathcal{H}, \mathcal{L})$.

In the sequel we shall say that $A$ is an increasing (decreasing) event if for all $(\mathcal{H}, \mathcal{L}) \in A$,

$$
\text { if }\left(\mathcal{H}^{\prime}, \mathcal{L}^{\prime}\right) \gg(\mathcal{H}, \mathcal{L})\left(\left(\mathcal{H}^{\prime}, \mathcal{L}^{\prime}\right) \ll(\mathcal{H}, \mathcal{L})\right) \text { then }\left(\mathcal{H}^{\prime}, \mathcal{L}^{\prime}\right) \in A
$$

As it has been proved in $[\mathrm{AKN}]$ a probability measure $\mathbb{P}_{N}^{\beta, \lambda}(\cdot)$ has the positive association property: if both $A$ and $B$ are increasing (decreasing) events, then

$$
\mathbb{P}_{N}^{\beta, \lambda}(A \cap B) \geqslant \mathbb{P}_{N}^{\beta, \lambda}(A) \mathbb{P}_{N}^{\beta, \lambda}(B)
$$

1.8. Structure of the paper. Our proof is built upon the classical treatment (see e.g. [BGJ]). An essential additional complication to be encountered is that in the quantum case two different clusters may share a spatial component without intersecting. In other words it can happen that there exists an index $i \in G_{N}$ and two disjoint clusters $\mathfrak{C}_{1}$ and $\mathfrak{C}_{2}$ such that $\mathfrak{C}_{l} \cap \mathbb{S}_{\beta}^{i} \neq \emptyset$ for $l=1,2$.

In Section 2 we set up most of the relevant notation and develop our basic inductive construction of percolation clusters. In the $\beta<\infty$ case the genealogical structure of percolation clusters could be ignored and, accordingly, our exposition could be slightly simplified. However, the multi-index notation we employ will become indispensable in the $\beta=\infty$ case [IL] and, besides, we feel that it gives a rather natural way to describe connected clusters.

Section 3 is devoted to the proofs of all our main results.

## 2. Construction of COnnected clusters

2.1. Underlying probability space. Let $\mathbf{M}_{+}$be the countable set of all finite multiindices $\underline{\alpha}$,

$$
\mathbf{M}_{+}=\left\{\underline{\alpha}=\left(\alpha_{1}, \ldots, \alpha_{n}\right) ; \alpha_{i} \in \mathbb{N}, n=1,2, \ldots\right\} .
$$

There is a total order on $\mathbf{M}_{+}$: we shall say that $\underline{\alpha} \prec \underline{\gamma}$ if either $\underline{\alpha}$ has less entries (belongs to an older generation) than $\underline{\gamma}$ or else if $\underline{\alpha}$ is less than $\underline{\gamma}$ in the lexicographical order. The underlying probability measure which we proceed to call $\mathbb{P}_{N}^{\beta, \lambda}(\cdot)$ is a product measure

$$
\mathbb{P}_{N}^{\beta, \lambda}(\cdot)=\bigotimes_{\underline{\alpha} \in \mathrm{M}_{+}} \mathbb{Q}_{N, \underline{\alpha}}^{\beta, \lambda} .
$$

Each measure $\mathbb{Q}_{N, \underline{\alpha}}^{\beta, \lambda}(\cdot)$ generates an interval $I[\underline{\alpha}] \subseteq \mathbb{S}_{\beta}$ and, subsequently, a point process $\mathcal{X}_{+}[\underline{\alpha}]$ on $\mathfrak{G}_{N}^{\beta}=\mathrm{G}_{N} \times \mathbb{S}_{\beta}$ according to the following procedure, whose relation to the background $(\mathcal{H}, \mathcal{L})$ process of holes and links should be obvious:

Construction of $I[\underline{\alpha}]$. Let $U, V$ be two independent $\operatorname{Exp}(\lambda)$ random variables. If $U+V \geqslant \beta$, then $I=\mathbb{S}_{\beta}$. Otherwise, $I=(-V, U) \subset \mathbb{S}_{\beta}$.

In the sequel we shall refer to the distribution of the interval $I[\underline{\alpha}]$, or more precisely to the distribution of its length $|I[\underline{\alpha}]|$, just constructed as to $\Gamma_{\beta}(2, \lambda)$ distribution. Obviously, in the $\beta=\infty$ case $|I|$ of $I$ is distributed as $\Gamma(2, \lambda)$ variable.

Construction of the number of offsprings and of the point process $\mathcal{X}_{+}[\underline{\alpha}]$. Given $I[\underline{\alpha}]$ and, in particular the length $|I[\underline{\alpha}]|$, sample the number of off-springs $\xi_{+}[\underline{\alpha}]$ from the Poisson distribution

$$
\xi_{+}[\underline{\alpha}] \sim \text { Poisson }(|I[\underline{\alpha}]|) .
$$

We shall denote the (unconditional) distribution of $\xi_{+}[\underline{\alpha}]$ constructed above as $\Xi_{\beta}(2, \lambda)$.
With $I[\underline{\alpha}]$ and $\xi_{+}[\underline{\alpha}]$ fixed, sample

$$
\mathcal{X}_{+}[\underline{\alpha}]=\left\{x_{1}[\underline{\alpha}], \ldots, x_{\xi_{+}}[\underline{\alpha}]\right\}=\left\{\left(d_{1}[\underline{\alpha}], \tau_{1}[\underline{\alpha}]\right), \ldots,\left(d_{\xi_{+}}[\underline{\alpha}], \tau_{\xi_{+}}[\underline{\alpha}]\right)\right\}
$$

where the departure times $\tau_{1}, \ldots, \tau_{\xi_{+}}$are i.i.d. random variables with the uniform distribution on $I[\underline{\alpha}]$, whereas the departure destinations $d_{1}, \ldots d_{\xi_{+}}$are i.i.d. uniform $\operatorname{Uni}\left(\mathrm{G}_{N}\right)$ random variables. For the latter use we define the set of all departure destinations from $\underline{\alpha}$,

$$
\mathcal{D}_{+}[\underline{\alpha}]=\operatorname{proj}_{G_{N}} \mathcal{X}_{+}[\underline{\alpha}] .
$$

2.2. Construction of $\mathfrak{C}(x)$. Let $x \in \mathfrak{G}_{N}^{\beta}$. The connected cluster $\mathfrak{C}(x)$ (see Subsection 1.2) is a disjoint union,

$$
\begin{equation*}
\mathfrak{C}(x)=\bigvee_{k} i_{k} \times J_{k} \tag{2.18}
\end{equation*}
$$

of intervals $J_{k} \subseteq \mathbb{S}_{\beta}$ with spatial coordinates $i_{k}$. These disjoint intervals will be labeled by a subset $\mathbf{M}^{x} \subset \mathbf{M}_{+}$of multi-indices. We shall always record the multi-indices from $\mathbf{M}^{x}$ in their increasing order, $\mathbf{M}^{x}=\left\{\underline{\alpha}_{1}, \underline{\alpha}_{2}, \ldots\right\}$. In this way we denote the spatial coordinate $i_{k}=i\left[\underline{\alpha}_{k}\right]$ and the associated interval $J_{k}=J\left[\underline{\alpha}_{k}\right]$ and rewrite (2.18) as

$$
\begin{equation*}
\mathfrak{C}(x)=\bigvee_{\underline{\alpha} \in \mathbf{M}^{x}} i[\underline{\alpha}] \times J[\underline{\alpha}] . \tag{2.19}
\end{equation*}
$$

Our construction of $\mathbf{M}^{x}$ and, accordingly, of $\{i[\underline{\alpha}], J[\underline{\alpha}]\}$ is an inductive one: At each stage we screen a certain multi-index $\underline{\alpha}$ and keep track of

- $\mathbf{M}^{x}(\underline{\alpha})$ - set of multi-indices which were already saturated into $\mathbf{M}^{x}$ before $\underline{\alpha}$.
- $\mathbf{R}^{x}(\underline{\alpha})$ - set of multi-indices (including $\underline{\alpha}$ itself) which are potential candidates for the membership in $\mathbf{M}^{x}$ and which are yet to be screened.

Both $\mathbf{M}^{x}$ and $\mathbf{R}^{x}$ are updated once $\underline{\alpha}$ is screened. The construction is complete whenever we finish an update with $\mathbf{R}^{x}=\emptyset$.

Notice that at each stage we also keep track both of space coordinates $i[\underline{\alpha}] \in G_{N}$ and of time coordinates $t[\underline{\alpha}] \in \mathbb{S}_{\beta}$ for all multi-indices $\underline{\alpha}$ from $\mathbf{M}^{x} \cup \mathbf{R}^{x}$. On the other hand, we sample intervals $J[\underline{\alpha}]$ and the associated point processes $\mathcal{X}[\underline{\alpha}] \subseteq \mathcal{X}_{+}[\underline{\alpha}]$ only at the moment when $\underline{\alpha}$ is screened.

The fact the construction below indeed reproduces the correct distribution of $\mathfrak{C}(x)$ is straightforward once we try to think about all the Poisson processes involved in terms of the usual approximation by Bernoulli trials.

Initial stage. For $x=(j, t)$ set

$$
\underline{\alpha}_{1}=(1) \quad i\left[\underline{\alpha}_{1}\right]=j \quad t\left[\underline{\alpha}_{1}\right]=t \quad \mathbf{M}^{x}=\emptyset \quad \mathbf{R}^{x}=\left\{\underline{\alpha}_{1}\right\}
$$

Screening stage. If $\mathbf{R}^{x}$ is empty, then stop. Otherwise, choose $\underline{\alpha}$ to be the minimal element of $\mathbf{R}^{x}$ (and set $\mathbf{R}^{x}(\underline{\alpha})=\mathbf{R}^{x}$ and $\mathbf{M}^{x}(\underline{\alpha})=\mathbf{M}^{x}$ ). There are two steps to be performed and several cases to be considered:
STEP 1 Deciding whether $\underline{\alpha}$ is to be included into $\mathbf{M}^{x}$ (Cases 2 and 3 below) and, if yes, sampling of $J[\underline{\alpha}]$.
CASE 1 If there exists $\underline{\gamma} \in \mathbf{M}^{x}(\underline{\alpha})$ (and then necessarily satisfying $\underline{\gamma} \prec \underline{\alpha}$ ) such that

$$
i[\underline{\alpha}]=i[\underline{\gamma}] \quad \text { and } \quad t[\underline{\alpha}] \in J[\underline{\gamma}],
$$

then remove $\underline{\alpha}$ from $\mathbf{R}^{x}$ and proceed to screen the next multi-index of $\mathbf{R}^{x}$.
CASE 2 If $i[\underline{\alpha}] \neq i[\underline{\gamma}]$ for every $\underline{\gamma} \in \mathbf{M}^{x}(\underline{\alpha})$, which means that

$$
\begin{equation*}
\mathfrak{C}_{\underline{\alpha}}(x) \triangleq \bigvee_{\underline{\gamma} \in \mathrm{M}^{x}(\underline{\alpha})} i[\underline{\gamma}] \times J[\underline{\gamma}] \tag{2.20}
\end{equation*}
$$

does not hit $\mathbb{S}_{\beta}^{i[\underline{\alpha}]}$, then set $J[\underline{\alpha}]=t[\underline{\alpha}]+I[\underline{\alpha}]$, where $I[\underline{\alpha}]$ is sampled from $\mathbb{P}_{N}^{\beta, \lambda}(\cdot)$ as described in Subsection 2.1.
CASE 3 In the remaining case,

$$
\mathfrak{C}_{\underline{\alpha}}(x) \cap \mathbb{S}_{\beta}^{i[\underline{\alpha}]} \neq \emptyset \quad \text { but } \quad(i[\underline{\alpha}], t[\underline{\alpha}]) \notin \mathfrak{C}_{\underline{\alpha}}(x)
$$

In such a situation define $J[\underline{\alpha}]$ as the connected component of $t[\underline{\alpha}]$ of

$$
(t[\underline{\alpha}]+I[\underline{\alpha}]) \backslash\left(\mathfrak{C}_{\underline{\alpha}}(x) \cap \mathbb{S}_{\beta}^{i[\underline{\alpha}]}\right)
$$

If either CASE 2 or CASE 3 took place then add $\underline{\alpha}$ to $\mathbf{M}^{x}$, remove $\underline{\alpha}$ from $\mathbf{R}^{x}$ and proceed with the second step.

STEP 2 Generating descendants of $\underline{\alpha}$. Sample $\xi_{+}[\underline{\alpha}]$ and, accordingly, the point process $\mathcal{X}_{+}[\underline{\alpha}]$ from the underlying distribution $\mathbb{P}_{N}^{\beta, \lambda}(\cdot)$. Screen all $k=1, \ldots, \xi_{+}$departures of $\mathcal{X}_{+}$ as follows:
CASE 1 If $t[\underline{\alpha}]+\tau_{k} \notin J[\underline{\alpha}]$, then ignore this $k$-th departure.
CASE 2 Otherwise register $k$-th departure as follows: Add $(\underline{\alpha}, k)$ to $\mathbf{R}^{x}$ and set,

$$
i[(\underline{\alpha}, k)]=d_{k} \quad \text { and } \quad t[(\underline{\alpha}, k)]=t[\underline{\alpha}]+\tau_{k}
$$

Return to the beginning of the screening stage.
In the sequel we shall use the following notation: For each $\underline{\alpha} \in \mathbf{M}^{x}$ we define $\mathbf{N}^{x}(\underline{\alpha})$ as the set of all registered descendants of $\underline{\alpha}$. The corresponding point process is $\mathcal{X}[\underline{\alpha}]=$
$\left\{(i[\underline{\gamma}], \tau[\underline{\gamma}]) ; \underline{\gamma} \in \mathbf{N}^{x}(\underline{\alpha})\right\}$. Finally, we shall denote the set of all spatial coordinates of registered descendants of $\underline{\alpha}$ as

$$
\mathcal{D}[\underline{\alpha}]=\operatorname{proj}_{G_{N}} \mathcal{X}[\underline{\alpha}] .
$$

2.3. The critical curve. As it becomes clear from the above construction of $\mathfrak{C}(x)$, the size of $\mathbf{M}^{x}$ is stochastically dominated by the total population size of Galton-Watson process with offspring distribution $\Xi_{\beta}[2, \lambda]$. Let $\xi \sim \Xi_{\beta}[2, \lambda]$. Evidently, $\xi$ has finite exponential moments. Furthermore,

$$
\mathbb{E}_{N}^{\beta, \lambda}(\xi)=N \mathbb{E}_{N}^{\beta, \lambda}\left(1-\mathrm{e}^{-|I| / N}\right)=O\left(\frac{1}{N}\right)+\mathbb{E}_{N}^{\beta, \lambda}(|I|)
$$

Now (with the usual convention $0 \cdot \infty=0$ in the $\beta=\infty$ case),

$$
\mathbb{E}_{N}^{\beta, \lambda}(|I|)=\mathbb{E}(U+V ; U+V<\beta)+\beta \mathbb{P}(U+V \geqslant \beta)
$$

where, as before, $U$ and $V$ are two independent exponential $\operatorname{Exp}(\lambda)$ random variables. Since $U+V \sim \Gamma(2, \lambda)$,

$$
\mathbb{P}(U+V \geqslant \beta)=\int_{\beta}^{\infty} \lambda^{2} t \mathrm{e}^{-\lambda t} \mathrm{~d} t=(\lambda \beta+1) \mathrm{e}^{-\lambda \beta}
$$

In the same fashion,

$$
\begin{aligned}
\mathbb{E}(U+V ; U+V \leqslant \beta) & =\int_{0}^{\beta} \lambda^{2} t^{2} \mathrm{e}^{-\lambda t} \mathrm{~d} t \\
& =\frac{2}{\lambda}\left(1-\mathrm{e}^{-\lambda \beta}\right)-\left(\beta^{2} \lambda+2 \beta\right) \mathrm{e}^{-\lambda \beta}
\end{aligned}
$$

Consequently,

$$
\begin{equation*}
\mathbb{E}_{N}^{\beta, \lambda}(|I|)=\frac{2}{\lambda}\left(1-\mathrm{e}^{-\lambda \beta}\right)-\beta \mathrm{e}^{-\lambda \beta}=F(\beta, \lambda) \tag{2.21}
\end{equation*}
$$

which is, of course, precisely the right hand side of (1.5).
2.4. LRO: size of $\mathcal{S}^{x}$. In the $L R O$-case of $F(\beta, \lambda)>1$ we shall confine the discussion to the case of finite $\beta<\infty$. Define

$$
\mathcal{S}^{x}=\operatorname{proj}_{\mathrm{G}_{N}} \mathfrak{C}(x)
$$

In other words $\mathcal{S}^{x}$ is the set of all different spatial coordinates of $\mathfrak{C}(x)$. Obviously,

$$
|\mathfrak{C}(x)| \leqslant \beta \# \mathcal{S}^{x}
$$

In fact, as we shall see in Section 3 a converse is also true in the sense that large size of $\# \mathcal{S}^{x}$ necessarily implies that $|\mathfrak{C}(x)|$ is also large. Meanwhile:
Lemma 2.1. In the LRO-case let $\delta>0$ be such that $(1-\delta) F(\beta, \lambda)>1$. Then, for every $\kappa>0$ there exists $c_{1}=c_{1}(\beta, \lambda, \kappa)$, such that

$$
\begin{equation*}
\mathbb{P}_{N}^{\beta, \lambda}\left(\# \mathcal{S}^{x} \in\left[c_{1} \log N, \delta N\right]\right)=o\left(\frac{1}{N^{\kappa}}\right) \tag{2.22}
\end{equation*}
$$

In the sequel we shall say that $\mathfrak{C}(x)$ is small if $\# \mathcal{S}^{x}<c_{1} \log N$ and, accordingly, that it is large if $\# \mathcal{S}^{x}>\delta N$.
Proof of Lemma 2.1. Let us go back to STEP 2 of the screening stage of Subsection 2.2. A multi-index descendant $\underline{\gamma} \in \mathbf{N}^{x}(\underline{\alpha})$ of $\underline{\alpha} \in \mathbf{M}^{x}$ is called free if its colour $i[\underline{\gamma}]$ is encountered for the first time in the course of our construction of $\mathfrak{C}(x)$. Formally, $\underline{\gamma}$ is free if

$$
i[\underline{\gamma}] \notin\left\{i[\underline{\nu}]: \underline{\nu} \prec \underline{\gamma} \quad \text { and } \quad \underline{\nu} \in \mathbf{M}^{x}(\underline{\alpha}) \cup \mathbf{R}^{x}(\underline{\alpha}) \cup \mathbf{N}^{x}(\underline{\alpha})\right\} .
$$

Notice that any free descendant of $\underline{\alpha}$ will be duly included into $\mathbf{M}^{x}$ at later screening stages. Also notice that the set $\mathbf{F}^{x} \subseteq \mathbf{M}^{x}$ of all free multi-indices is in one-to-one correspondence with $\mathcal{S}^{x}$. Let us define

$$
\mathcal{S}^{x}(\underline{\alpha})=\left\{i[\underline{\gamma}]: \underline{\gamma} \in \mathbf{M}^{x}(\underline{\alpha}) \cup \mathbf{R}^{x}(\underline{\alpha})\right\} .
$$

In other words $\mathcal{S}^{x}(\underline{\alpha})$ is the set of all different spatial coordinates of $\mathfrak{C}(x)$ which were generated before screening of $\underline{\alpha}$. Then the number $\eta_{f}(\underline{\alpha})$ of free descendants of $\underline{\alpha}$ is given by

$$
\begin{equation*}
\eta_{f}(\underline{\alpha})=\sum_{i \in \mathrm{G}_{N} \backslash \mathcal{S}^{x}(\underline{\alpha})} \mathbb{I}_{i \in \mathcal{D}[\underline{\alpha}]} . \tag{2.23}
\end{equation*}
$$

If $\underline{\alpha}$ is itself free and, moreover, $\# \mathcal{S}^{x}(\underline{\alpha})<\delta N$, then $\eta_{f}(\underline{\alpha})$ stochastically dominates and, in fact, could be coupled in a straightforward way with a random variable $\eta[\underline{\alpha}]$ whose conditional on $I[\underline{\alpha}]$ distribution is binomial $\operatorname{Bin}\left((1-\delta) N, 1-\mathrm{e}^{-\mid I[\underline{\alpha}] / N}\right)$. Of course, $\mathbb{E}_{N}^{\beta, \lambda}(\eta) \rightarrow(1-\delta) F(\beta, \lambda)>1$ as $N \rightarrow \infty$.

Let us, thereby, consider an i.i.d. family $\left\{\eta_{k}\right\}$ of random variables distributed as above, which is coupled with the collection $\left\{\eta_{f}(\underline{\alpha})\right\}$ in the following way: Write all free multiindices $\mathbf{F}^{x}=\left\{\underline{\alpha}_{1}^{*}, \underline{\alpha}_{2}^{*}, \ldots\right\}$ in their increasing order. Then $\eta_{k} \leqslant \eta_{f}\left[\underline{\alpha}_{k}^{*}\right]$ on the event $\left\{\# \mathcal{S}^{x}\left(\underline{\alpha}_{k}^{*}\right)<\delta N\right\}$. Since $\# \mathcal{S}^{x}=\# \mathbf{F}^{x}$, and

$$
\begin{equation*}
\# \mathcal{S}^{x} \geqslant \sum_{\underline{\alpha} \in \mathbf{F}^{x}} \eta_{f}(\underline{\alpha}) \tag{2.24}
\end{equation*}
$$

we infer that

$$
\mathbb{P}_{N}^{\beta, \lambda}\left(\# \mathcal{S}^{x} \in\left[c_{1} \log N, \delta N\right]\right) \leqslant \sum_{n=c_{1} \log N}^{\delta N} \mathbb{P}_{N}^{\beta, \lambda}\left(\sum_{1}^{n} \eta_{k} \leqslant n\right)
$$

and (2.22) follows from a usual LD upper bound, as $\mathbb{E}_{N}^{\beta, \lambda}(\eta)>1$.
2.5. Coupling between two clusters. Recall that we are recording the multi-indices $\mathbf{M}^{x}=\left\{\underline{\alpha}_{1}, \underline{\alpha}_{2}, \ldots\right\}$ according to the order in which they were saturated or, equivalently, in their increasing order. With each $k=1,2, \ldots$ we associate a $\sigma$-algebra

$$
\begin{equation*}
\Sigma_{k}^{x}=\sigma\left(\mathcal{X}_{+}\left[\underline{\alpha}_{l}\right], I\left[\underline{\alpha}_{l}\right] ; l \leqslant k\right\} . \tag{2.25}
\end{equation*}
$$

In order to avoid ambiguities we set $\underline{\alpha}_{k} \equiv \underline{\alpha}_{\mathfrak{m}^{x}}$ for $k \geqslant \mathfrak{m}^{x}$, where $\mathfrak{m}^{x} \triangleq \# \mathbf{M}^{x}$. Clearly, $\Sigma_{k}^{x}$ contains all the information on the growth of $\mathfrak{C}(x)$ up to the $k$ th (or, in the case $k>\mathfrak{m}^{x}$, up to the last) saturation. In particular, the eventual number of intervals $\mathfrak{m}^{x}$ in the decomposition (2.18) is a stopping time with respect to the filtration $\left\{\Sigma_{k}^{x}\right\}$.

Our proof of various claims in the LRO case is based on the following modification of the growth algorithm described in Subsection 2.2: Let $x, y$ be two points of $\mathfrak{G}_{N}^{\beta}$.

First of all we shall grow $\mathfrak{C}(x)$ according to the rules of Subsection 2.2 until a certain stopping time $S$ (with respect to the filtration $\left\{\Sigma_{k}^{x}\right\}$ ). We shall define $S$ below. In any case, the relevant notation is (see (2.20)),

$$
\mathfrak{C}_{k}(x)=\mathfrak{C}_{\underline{\alpha}_{k+1}}(x), \mathbf{M}_{k}^{x}=\mathbf{M}^{x}\left(\underline{\alpha}_{k+1}\right) \quad \text { and } \quad \mathbf{R}_{k}^{x}=\mathbf{R}^{x}\left(\underline{\alpha}_{k+1}\right) .
$$

Accordingly, $\mathfrak{C}_{S}(x)$ is the shape of the cluster whose growth was suspended at $S$, and $\mathbf{R}_{S}^{x}$ is the set of all multi-indices which were registered but not yet saturated (or dropped) by time $S$.

Secondly, we shall grow $\mathfrak{C}(y)$ according to the following modified set of rules, which should necessarily take into account already existing information as recorded in $\Sigma_{S}^{x}$.

Growth of $\mathfrak{C}(y)$ will be halted at some random stage and certainly once it will become clear that $\mathfrak{C}(x)$ and $\mathfrak{C}(y)$ are going to merge. In the sequel we shall use $T$ to denote the stage at which we halt the construction of $\mathfrak{C}(y)$. We use $\tilde{\mathbf{M}}^{y}, \tilde{\mathbf{R}}^{y}$ etc notation to stress that we construct $\mathfrak{C}(y)$ after all the data for $\Sigma_{S}^{x}$ has been already sampled. If $\tilde{\mathbf{M}}^{y}=\left\{\underline{\gamma}_{1}, \underline{\gamma}_{2}, \ldots\right\}$ is the (ordered) set of all saturated multi-indices of $\mathfrak{C}(y)$, then $T$ will be a the stopping time with respect to the filtration

$$
\tilde{\Sigma}_{0}^{y}=\Sigma_{S}^{x} \quad \text { and } \quad \tilde{\Sigma}_{k}^{y}=\sigma\left(\Sigma_{S}^{x}, \Sigma_{k}^{y}\right)
$$

where $\Sigma_{k}^{y}$ is defined as in (2.25).

The initial stage. If $y \in \mathfrak{C}_{S}(x)$, then set $T=0$ and halt immediately. Otherwise proceed as in the initial stage of one-cluster construction of Subsection 2.2 (except for setting $\left(\underline{\gamma}_{1}\right)=(2)$, so that $x$ and $y$ related multi-indices are distinguished).

An update stage. There are several corrections which are due to the already existing information contained in $\Sigma_{S}^{x}$. Let $\underline{\gamma} \in \tilde{\mathbf{R}}^{y}$ be a multi-index under screening.
CORRECTION 1 In STEP 1 if $(i[\underline{\gamma}], t[\underline{\gamma}]) \in \mathfrak{C}_{S}(x)$, then remove $\underline{\gamma}$ from $\tilde{\mathbf{R}}^{y}$ and proceed to screen the next multi-index of $\tilde{\mathbf{R}}^{\bar{y}}$.
CORRECTION 2 In CASE 2 and CASE 3 of STEP 1 define the interval $J[\underline{\gamma}]$ as the connected component of $t[\underline{\gamma}]$ of

$$
\left((t[\underline{\gamma}]+I[\underline{\gamma}]) \backslash\left(\mathfrak{C}_{\underline{\gamma}}(y) \cap \mathfrak{C}_{S}(x) \cap \mathbb{S}_{\beta}^{i[\underline{\gamma}]}\right)\right.
$$

CORRECTION 3 If, after $J[\underline{\gamma}]$ is sampled, there exists $\underline{\alpha} \in \mathbf{R}_{S}^{x}$ such that $i[\underline{\alpha}]=i[\underline{\gamma}]$ and $t[\underline{\alpha}] \in J[\underline{\gamma}]$, then the clusters of $x$ and $y$ are going to merge. If $\underline{\gamma}=\underline{\gamma}_{k}$, then set $T=\bar{k}$ and halt the construction of $\mathfrak{C}(y)$.
2.6. Merging of two large clusters. Recall that a cluster $\mathfrak{C}(x)$ is called large if $\#\left(\mathcal{S}^{x}\right)>$ $\delta N$, where $\delta>0$ is some fixed number satisfying $(1-\delta) F(\beta, \lambda)>1$.

Lemma 2.2. Let $\beta<\infty$. Then for every $\kappa>0$,

$$
\begin{equation*}
\mathbb{P}_{N}^{\beta, \lambda}(\mathfrak{C}(x) \cap \mathfrak{C}(y)=\emptyset \mid \text { both } \mathfrak{C}(x) \text { and } \mathfrak{C}(y) \text { are large })=o\left(\frac{1}{N^{\kappa}}\right) \tag{2.26}
\end{equation*}
$$

uniformly in $x, y \in \mathfrak{G}_{N}^{\beta}$.
Proof of Lemma 2.2 The proof relies on the construction of clusters and, accordingly, on the coupling between $\mathfrak{C}(x)$ and $\mathfrak{C}(y)$ as introduced in Subsection 2.2 and Subsection 2.5.
STEP 1 We start by growing the cluster $\mathfrak{C}(x)$ until the stopping time $S$ which is described as follows: Let $\epsilon>0$ be small, e.g. $3 \epsilon<\delta$ would suffice. Then,

$$
\begin{equation*}
S=\max \left\{k \leqslant \mathfrak{m}^{x}: \# \mathcal{S}^{x}\left(\underline{\alpha}_{k}\right)<\epsilon N\right\} \tag{2.27}
\end{equation*}
$$

Evidently, $S$ is a stopping time with respect to the filtration $\left\{\Sigma_{k}^{x}\right\}$.
Let $\mathbf{F}_{S}^{x}$ be the set of all free multi-indices of $\mathfrak{C}(x)$ which were generated by time $S$. In this notation,

$$
\mathcal{S}^{x}\left(\underline{\alpha}_{S}\right) \cup \mathcal{D}\left(\underline{\alpha}_{S}\right)=\bigcup_{\underline{\alpha} \in \mathbf{F}_{S}^{x}} i[\underline{\alpha}] .
$$

We shall distinguish between the free multi-indices $\mathbf{F}_{S,-}^{x}=\mathbf{F}_{S}^{x} \cap \mathbf{M}_{S}^{x}$ which were already saturated by time $S$ and the remaining free multi-indices $\mathbf{F}_{S,+}^{x}=\mathbf{F}_{S}^{x} \backslash \mathbf{F}_{S,-}^{x}$. Let $\nu>0$ be such that

$$
\begin{equation*}
(1-3 \nu)(1-\delta) F(\beta, \lambda)>1 \tag{2.28}
\end{equation*}
$$

We claim that once $\mathfrak{C}(x)$ is large, then up to $o\left(\frac{1}{N^{\kappa}}\right)$-probabilities the cardinality of unsaturated free multi-indices $\# \mathbf{F}_{S,+}^{x}$ exceeds $3 \nu \epsilon N$, or, more precisely, the cardinality of saturated multi-indices $\# \mathbf{F}_{S,+}^{x}$ does not exceed $(1-3 \nu) \epsilon N$. Indeed, by construction $\# \mathbf{F}_{S}^{x}>\epsilon N$, whereas $\# \mathcal{S}^{x}\left(\underline{\alpha}_{S}\right)<\epsilon N$. On the other hand, as in (2.24),

$$
\# \mathcal{S}^{x}\left(\underline{\alpha}_{S}\right) \geqslant \sum_{\underline{\alpha}_{k}^{*} \in \mathbf{F}_{S,-}^{x} \backslash \underline{\alpha}_{S}} \eta_{k}
$$

Consequently,

$$
\begin{align*}
\mathbb{P}_{N}^{\beta, \lambda}\left(\# \mathbf{F}_{S,+}^{x}<3 \nu \epsilon N ; \mathfrak{C}(x) \text { is large }\right) \leqslant & \mathbb{P}_{N}^{\beta, \lambda}\left(\# \mathbf{F}_{S,-}^{x}>(1-3 \nu) \epsilon N ; \#\left(\mathcal{S}^{x}\left(\underline{\alpha}_{S}\right)\right)<\epsilon N\right) \\
& \leqslant \mathbb{P}_{N}^{\beta, \lambda}\left(\sum_{1}^{(1-3 \nu) \epsilon N} \eta_{k}<\epsilon N\right)=o\left(\frac{1}{N^{\kappa}}\right) \tag{2.29}
\end{align*}
$$

since $(1-3 \nu) \mathbb{E}_{N}^{\beta, \lambda}\left(\eta_{k}\right)$ is still larger than 1 whenever $\nu$ complies with (2.28).
STEP 2 We shall now switch to the coupled construction of $\mathfrak{C}(y)$. First of all we shall adjust the notion of the set $\tilde{\mathbf{F}}^{y}$ of free multi-indices of $\mathfrak{C}(y)$ : Recall that free multi-indices are generated at STEP 2 of the screening stage when we consider family of multi-index descendants $\mathbf{N}^{y}(\underline{\alpha})$ of freshly saturated multi-indices $\underline{\alpha} \in \tilde{\mathbf{M}}^{y}$. We shall say that $\underline{\gamma} \in$ $\mathbf{N}^{y}(\underline{\alpha})$ is free if, as before, its colour $i[\gamma]$ is encountered for the first time in the course of construction of $\mathfrak{C}(y)$, but, in addition, $i[\underline{\gamma}]$ was not saturated in $\mathfrak{C}_{S}(x)$, that is:

$$
i[\underline{\gamma}] \notin\left\{i[\underline{\nu}]: \underline{\nu} \in \mathbf{M}_{S}^{x}\right\} .
$$

The latter set is in one-to-one correspondence with $\mathbf{F}_{S,-}^{x}$. As a result, on the event $\{\mathfrak{C}(y)$ is large and $\mathfrak{C}(y) \cap \mathfrak{C}(x)=\emptyset\}$,

$$
\# \tilde{\mathbf{F}^{y}} \geqslant \# \mathcal{S}^{y}-\# \mathbf{F}_{S,-}^{x} \geqslant \delta N-(1-3 \nu) \epsilon N \geqslant 2 \epsilon N
$$

where the first inequality (up to $o\left(\frac{1}{N^{\kappa}}\right)$-probability) follows from (2.29) and the second inequality follows from our choice of $3 \epsilon<\delta$. Define now a $\left\{\tilde{\Sigma}_{k}^{y}\right\}$-stopping time

$$
T=\min \left\{k: \# \tilde{\mathbf{F}}^{y}\left(\underline{\gamma}_{k}\right) \geqslant \nu \epsilon N\right\}
$$

where, $\tilde{\mathbf{F}}^{y}\left(\underline{\gamma}_{k}\right)$ is the set of all (modified) free multi-indices which were generated during the first $k$ screening sessions. Since by construction (end of Subsection 2.1) $\left\{\# \mathcal{D}_{+}[\underline{\alpha}]\right\}$ is a family of i.i.d random variables with exponentially decaying tails, we can safely assume that, up to $o\left(\frac{1}{N^{\kappa}}\right)$-probability

$$
\begin{equation*}
\# \tilde{\mathbf{F}}^{y}\left(\underline{\gamma}_{T}\right) \leqslant 2 \nu \epsilon N \tag{2.30}
\end{equation*}
$$

Similarly to (2.29), set $\tilde{\mathbf{F}}_{T,-}^{y}=\tilde{\mathbf{M}}_{T}^{y} \cap \tilde{\mathbf{F}}_{T}^{y}$. Then,

$$
\mathbb{P}_{N}^{\beta, \lambda}\left(\# \tilde{\mathbf{F}}_{T,-}^{y}>(1-3 \epsilon) \# \tilde{\mathbf{F}}_{T}^{y} ; \# \tilde{\mathbf{F}}_{T}^{y} \in[\nu \epsilon N, 2 \nu \epsilon N]\right)=o\left(\frac{1}{N^{\kappa}}\right)
$$

Consequently, up to $o\left(\frac{1}{N^{\kappa}}\right)$-probability, at least $3 \nu \epsilon^{2} N$ modified free multi-indices from $\tilde{\mathbf{F}}_{T}^{y}$ were not saturated by time $T$. Let us denote the latter set of multi-indices as $\tilde{\mathbf{F}}_{S, T}^{y}$,

$$
\begin{equation*}
\# \tilde{\mathbf{F}}_{S, T}^{y} \geqslant 3 \nu \epsilon^{2} N \tag{2.31}
\end{equation*}
$$

On the other hand, the number of all saturated multi-indices from $\tilde{\mathbf{F}}_{T}^{y}$ is, in view of (2.30), bounded above by $2 \nu \epsilon N$, which falls short of at least $3 \nu \epsilon$ free unsaturated multi-indices of $\mathbf{F}_{S,+}^{x}$. Thus the cardinality of the set

$$
\tilde{\mathbf{F}}_{S, T}^{x} \triangleq\left\{\underline{\alpha} \in \mathbf{F}_{S,+}^{x}: i[\underline{\alpha}] \notin \operatorname{proj}_{G_{N}} \mathfrak{C}_{T}(y)\right\}
$$

satisfies,

$$
\begin{equation*}
\# \tilde{\mathbf{F}}_{S, T}^{x} \geqslant \nu \epsilon N \tag{2.32}
\end{equation*}
$$

STEP 3 In view of (2.31) and (2.32) let us summarize the above results as follows: Consider the event $\{\mathfrak{C}(x) \cap \mathfrak{C}(y)=\emptyset$; both $\mathfrak{C}(x)$ and $\mathfrak{C}(y)$ are large $\}$ and let $S$ and $T$ be stopping times defined as above. Then, up to $o\left(\frac{1}{N^{\kappa}}\right)$-probability, there exists a set of unsaturated free multi-indices $\tilde{\mathbf{F}}_{S, T}^{x}$ and, accordingly, a set of unsaturated free multi-indices $\tilde{\mathbf{F}}_{S, T}^{y}$ such that,
(1) $\# \tilde{\mathbf{F}}_{S, T}^{x} \geqslant \nu \epsilon N$ and $\# \tilde{\mathbf{F}}_{S, T}^{y} \geqslant 3 \nu \epsilon^{2} N$.
(2) For every $\underline{\alpha} \in \tilde{\mathbf{F}}_{S, T}^{x} \cup \tilde{\mathbf{F}}_{S, T}^{y}$ and every $\underline{\gamma} \in \mathbf{M}_{S}^{x} \cup \tilde{\mathbf{M}}_{T}^{y}$ the colours $i[\underline{\alpha}] \neq i[\underline{\gamma}]$.
(3) Any two different multi-indices of $\tilde{\mathbf{F}}_{S, T}^{x}$, respectively of $\tilde{\mathbf{F}}_{S, T}^{y}$, have different colours. Set

$$
\mathcal{S}_{S, T}^{x, y}=\left\{i[\underline{\alpha}]: \underline{\alpha} \in \tilde{\mathbf{F}}_{S, T}^{x} \cup \tilde{\mathbf{F}}_{S, T}^{y}\right\} .
$$

Property (2) above means that none of the processes of holes $\mathcal{H}_{i}$ was ever sampled for $i \in \mathcal{S}_{S, T}^{x, y}$. Similarly none of the processes $\mathcal{L}_{i j}$ of links was ever sampled for $i, j \in \mathcal{S}_{S, T}^{x, y}$. Since the family of processes

$$
\left\{\mathcal{H}_{i}, \mathcal{L}_{i j}\right\}_{i, j \in \mathcal{S}_{S, T}^{x, y}}
$$

is, conditionally on $\mathcal{S}_{S, T}^{x, y}$, independent of $\tilde{\Sigma}_{T}^{y}$, we infer that up to a $o\left(\frac{1}{N^{\kappa}}\right)$-probability,

$$
\begin{equation*}
\mathbb{P}_{N}^{\beta, \lambda}\left(\mathfrak{C}(x) \nleftarrow \mathfrak{C}(y) \mid \tilde{\Sigma}_{T}^{y}\right) \leqslant \max _{\substack{v_{1}, \ldots, v_{L} \\ u_{1}, \ldots, u_{M}}} \mathbb{P}_{N}^{\beta, \lambda}\left(\bigcap_{i, j}\left\{I\left(v_{i}\right) \nleftarrow I\left(u_{j}\right)\right\}\right) \tag{2.33}
\end{equation*}
$$

where $L=\nu \epsilon N, M=3 \nu^{2} \epsilon N$ and the maximum above is over all colour disjoint collections (see property (3) above) of points $\left\{v_{1}, \ldots, v_{L}\right\} \subset \mathfrak{G}_{N}^{\beta}$ and $\left\{u_{1}, \ldots, u_{M}\right\} \subset \mathfrak{G}_{N}^{\beta}$. We have used the obvious notation in (2.33): For $z=(i, t) \in \mathfrak{G}_{N}^{\beta}, I(z)$ denotes the interval which contains $t$ in the decomposition (1.3) of $\mathbb{S}_{\beta}^{i}$.
STEP 4 It remains to derive an upper bound on the right-hand side of (2.33). Set $K=$ $L-M=\nu(1-3 \nu) \epsilon N$. By our choice of $\nu$ in (2.28) the number $K$ is positive and, moreover, proportional to $N$. At least $K$ of colour disjoint $u_{j}$-s have spatial coordinates different from any of $u_{i}$ spatial coordinates. There is no loss to assume that this property is enjoyed by the first $K$ points $\left\{v_{1}, \ldots, v_{K}\right\}$. Thus, the right hand side of (2.33) is bounded above by

$$
\begin{equation*}
\max _{\substack{v_{1}, \ldots, v_{K} \\ u_{1}, \ldots, u_{M}}} \mathbb{P}_{N}^{\beta, \lambda}\left(\bigcap_{i, j}\left\{I\left(v_{i}\right) \nleftarrow I\left(u_{j}\right)\right\}\right), \tag{2.34}
\end{equation*}
$$

where the maximum is now over all possible $K+L$ colour disjoint points of $\mathfrak{G}_{N}^{\beta}$. By usual large deviation estimates there is a constant $c=c(\beta \epsilon, \nu)>0$ such that up to $\mathrm{e}^{-c N_{-}}$ probability there are at least $\mathrm{e}^{-\lambda \beta} K / 2$ of $v_{j}$-s lie on $\mathbb{S}_{\beta}$ circles without holes, and at least $\mathrm{e}^{-\lambda \beta} M / 2$ of $u_{j}$-s which also lie on $\mathbb{S}_{\beta}$ circles without holes. This is a reduction to the computation employed in the classical case: Indeed, for any two $v_{i}$ and $u_{j}$ with spatial coordinates $i \neq j$,

$$
\mathbb{P}_{N}^{\beta, \lambda}\left(I\left(v_{i}\right) \nleftarrow I\left(u_{j}\right) \mid \mathcal{H}_{i}=\emptyset, \mathcal{H}_{j}=\emptyset\right)=\mathrm{e}^{-\beta / N} .
$$

Consequently, the expression in (2.34) is bounded above by

$$
\mathrm{e}^{-c N}+\exp \left\{-\frac{\beta \mathrm{e}^{-2 \lambda \beta} K M}{4 N}\right\}
$$

Since $K M / N=3 \nu^{2}(1-3 \nu) \epsilon^{2} N$, the proof of Lemma 2.2 is complete.

## 3. Proofs of main results

3.1. Short range order. We employ the construction and the notation of Subsection 2.1 and Subsection 2.2.

Proof of (1.7) and of (1.11) for $\beta<\infty$. . Let $\mathcal{Z}_{+}$be the sub-critical Galton-Watson process with offspring distribution $\Xi_{\beta}(2, \lambda)$ (Subsection 2.1). The members of $\mathcal{Z}_{+}$are naturally labeled by multi-indices from $\mathbf{M}_{+}$and, furthermore, it is straightforward to couple the construction of $\mathcal{Z}_{+}$with that of $\mathfrak{C}(x)$ in such a way that

$$
\mathbf{M}^{x} \subseteq \mathcal{Z}_{+} \quad \text { and } \quad|\mathfrak{C}(x)| \leqslant \sum_{\underline{\alpha} \in \mathcal{Z}_{+}}|I[\underline{\alpha}]| \triangleq\left|\mathcal{Z}_{+}\right|
$$

As a result for every $\kappa>0$ there exists $c>0$ such that

$$
\begin{equation*}
\mathbb{P}_{N}^{\beta, \lambda}\left(\#\left(\mathcal{S}^{x}\right)>c \log N\right) \leqslant \mathbb{P}_{N}^{\beta, \lambda}\left(\# \mathcal{Z}_{+}>c \log N\right)=o\left(\frac{1}{N^{\kappa}}\right) \tag{3.35}
\end{equation*}
$$

Since, $|\mathfrak{C}(x)| \leqslant \beta \# \mathcal{S}^{x}$, the bound (1.11) readily follows in the $\beta<\infty$ case.
Furthermore, if $x=(i, t)$ and $j \neq i$, then by the $\mathrm{G}_{N}$-permutation invariance,

$$
\begin{equation*}
\mathbb{P}_{N}^{\beta, \lambda}\left(j \in \mathcal{S}^{x} \mid \# \mathcal{S}^{x}=n\right)=\frac{n-1}{N-1} \tag{3.36}
\end{equation*}
$$

which, in view of (3.35), implies the SRO bound (1.7) at any $\beta \leqslant \infty$.
Proof of (1.11) in the $\beta=\infty$ case. First of all notice that in the $\beta=\infty$ case, the underlying distribution $\Xi_{\infty}(2, \lambda)$ of offsprings is just negative binomial (for the number of failures),

$$
\begin{equation*}
\mathbb{P}\left(\xi_{+}=k\right)=(k+1) \frac{\lambda^{2}}{(1+\lambda)^{k+2}} \quad k=0,1,2, \ldots, \tag{3.37}
\end{equation*}
$$

whereas the conditional distribution of the interval length $|I|$ is

$$
\begin{equation*}
|I| \sim \Gamma(k+2,1+\lambda) \quad \text { given } \xi_{+}=k \text { for } k=0,1,2, \ldots \tag{3.38}
\end{equation*}
$$

In the $\beta=\infty$ case formula (3.38) implies that the conditional distribution of $\left|\mathcal{Z}_{+}\right|$given the value of the total size of the population $\# \mathcal{Z}_{+}$is precisely that of the sum of $3 \# \mathcal{Z}_{+}-1$
independent exponential $\operatorname{Exp}(1+\lambda)$ random variables. Therefore,

$$
\begin{align*}
\mathbb{P}_{N}^{\beta, \lambda}\left(\left|\mathcal{Z}_{+}\right|>c \log N\right) \leqslant & \mathbb{P}_{N}^{\beta, \lambda}\left(\# \mathcal{Z}_{+}>\frac{1+\lambda}{6} c \log N\right) \\
& +\max _{n \leqslant \frac{c(1+\lambda) \log N}{6}} \mathbb{P}_{N}^{\beta, \lambda}\left(\sum_{1}^{3 n}\left(\eta_{i}-\mathbb{E}_{N}^{\beta, \lambda}\left(\eta_{i}\right)\right)>\frac{c}{2} \log N\right) \tag{3.39}
\end{align*}
$$

where $\eta_{1}, \eta_{2}, \ldots$ are i.i.d. $\operatorname{Exp}(1+\lambda)$ random variables. Clearly, for every $\kappa>0$ one can choose $c$ sufficiently large so that the right-hand side of (3.39) is bounded above by $1 / N^{\kappa}$.

Proof of (1.12). Let $\beta<\infty$. By the union bound and $\mathrm{G}_{N}$-permutation invariance of the distribution of $\mathcal{M}$,

$$
\mathbb{P}_{N}^{\beta, \lambda}(\mathcal{M} \geqslant c \log N) \leqslant N \mathbb{P}_{N}^{\beta, \lambda}\left(\mathcal{M} \geqslant c \log N ; \mathfrak{C}^{*} \cap \mathbb{S}_{\beta}^{i} \neq \emptyset\right)
$$

where $\mathfrak{C}^{*}$ is the maximal cluster in the decomposition (1.2), $\left|\mathfrak{C}^{*}\right|=\mathcal{M}$. At this stage define,

$$
\mathcal{M}_{i}=\max \left\{\left|\mathfrak{C}_{k}\right|: \mathfrak{C}_{k} \cap \mathbb{S}_{\beta}^{i} \neq \emptyset\right\}
$$

Obviously,

$$
\left\{\mathcal{M} \geqslant c \log N ; \mathfrak{C}^{*} \cap \mathbb{S}_{\beta}^{i} \neq \emptyset\right\} \subseteq\left\{\mathcal{M}_{i} \geqslant c \log N\right\}
$$

Since the latter event is already increasing,

$$
\mathbb{P}_{N}^{\beta, \lambda}\left(\mathcal{M}_{i}>c \log N\right) \leqslant \mathbb{P}_{N}^{\beta, \lambda}\left(\mathcal{M}_{i}>c \log N \mid \mathcal{H}_{i}=\emptyset\right)
$$

However,

$$
\mathbb{P}_{N}^{\beta, \lambda}\left(\mathcal{M}_{i}>c \log N \mid \mathcal{H}_{i}=\emptyset\right) \leqslant \frac{\mathbb{P}_{N}^{\beta, \lambda}(|\mathfrak{C}((i, 0))|>c \log N)}{\mathbb{P}_{N}^{\beta, \lambda}\left(\mathcal{H}_{i}=\emptyset\right)}
$$

Since $\mathbb{P}_{N}^{\beta, \lambda}\left(\mathcal{H}_{i}=\emptyset\right)=\mathrm{e}^{-\lambda \beta}$, the target bound (1.12) follows.
3.2. LRO: Proof of Theorem B. For the rest of the paper we shall assume that $\beta<\infty$ and that $F(\beta, \lambda)>1$.

Probability of large cluster. Recall that we say that $\mathfrak{C}(x)$ is large if the number of spatial coordinates satisfies $\# \mathcal{S}^{x} \geqslant \delta N$, where $\delta$ is a fixed small number, $(1-\delta) F(\beta, \lambda)>1$.

Both upper and lower bounds on $\mathbb{P}_{N}^{\beta, \lambda}(\mathfrak{C}(x)$ is large) follow, as in the classical case, from comparison with appropriate Galton-Watson processes.

The upper bound is straightforward: As in Subsection 3.1 let $\mathcal{Z}_{+}$be the Galton-Watson process with offspring distribution $\Xi_{\beta}(2, \lambda)$. Define $p(\beta, \lambda)$ as the survival probability of $\mathcal{Z}_{+}$, that is $p(\beta, \lambda)$ is the unique non-trivial root of

$$
\begin{equation*}
1-p=\mathbb{E}_{N}^{\beta, \lambda}\left((1-p)^{\xi_{+}}\right) \tag{3.40}
\end{equation*}
$$

where $\xi_{+} \sim \Xi_{\beta}(2, \lambda)$. By the very construction of $\mathfrak{C}(x)$ in Subsection 2.2,

$$
\mathbb{P}_{N}^{\beta, \lambda}(\mathfrak{C}(x) \text { is large }) \leqslant p(\beta, \lambda)(1+o(1))
$$

The lower bound is slightly more delicate. It relies on a coupling with yet another GaltonWatson process $\mathcal{Z}_{+}^{f}$, which lives on free multi-indices $\mathbf{F}^{x}$ of $\mathfrak{C}(x)$. Specifically, as before set $\mathbf{F}^{x}=\left\{\underline{\alpha}_{1}^{*}, \underline{\alpha}_{2}^{*}, \ldots\right\}$. Recall the definition of random variables $\eta_{f}\left(\underline{\alpha}_{k}^{*}\right)$ in (2.23). On the event

$$
\left\{\# \mathcal{S}^{x}<c \log N\right\}=\left\{\# \mathbf{F}^{x}<c \log N\right\}
$$

the sequence $\left\{\eta_{f}\left(\underline{\alpha}_{k}^{*}\right)\right\}$ can be coupled with an i.i.d sequence $\left\{\eta_{k}^{*}\right\}$, such that

$$
\forall \underline{\alpha}_{k}^{*} \in \mathbf{F}^{x} \quad \eta_{f}\left(\underline{\alpha}_{k}^{*}\right) \geqslant \eta_{k}^{*}
$$

whereas the distribution of $\eta^{*}$ is given by: First sample $|I|$ from $\Gamma_{\beta}(2, \lambda)$, and then sample $\eta^{*}$ from $\operatorname{Bin}\left(N-c \log N, 1-\mathrm{e}^{-|I| / N}\right)$.

Let $p_{N}(\beta, \lambda)$ be the survival probability of the Galton-Watson process with the offspring distribution specified by $\eta$ above. Clearly,

$$
\begin{equation*}
\lim _{N \rightarrow \infty} p_{N}(\beta, \lambda)=p(\beta, \lambda) \tag{3.41}
\end{equation*}
$$

On the other hand, by Lemma 2.1,

$$
\mathbb{P}_{N}^{\beta, \lambda}(\mathfrak{C}(x) \text { is large })=1-\mathbb{P}_{N}^{\beta, \lambda}\left(\# \mathcal{S}^{x} \leqslant c \log N\right) \geqslant 1-\mathbb{P}_{N}^{\beta, \lambda}\left(\mathcal{Z}_{+}^{f} \text { dies out }\right)
$$

up to $o(1)$ probabilities. In view of (3.41) this gives a complementary lower bound and, consequently,

$$
\begin{equation*}
\mathbb{P}_{N}^{\beta, \lambda}(\mathfrak{C}(x) \text { is large })=p(\beta, \lambda)(1+o(1)) \tag{3.42}
\end{equation*}
$$

Size of $\mathcal{M}$. It is a straightforward exercise to deduce from (3.42), Lemma 2.1 and from the FKG properties of $\mathbb{P}_{N}^{\beta, \lambda}(\cdot)$ that in the LRO regime

$$
\mathbb{P}_{N}^{\beta, \lambda}\left(\exists x \in \mathfrak{G}_{N}^{\beta} \text { s.t. } \mathfrak{C}(x) \text { is large }\right)=1-O\left(\frac{1}{N^{\kappa}}\right)
$$

for every $\kappa>0$. Furthermore, it is equally straightforward to deduce from Lemma 2.2 that

$$
\mathbb{P}_{N}^{\beta, \lambda}(\exists \text { unique large cluster } \mid \exists \text { large cluster })=1-O\left(\frac{1}{N^{\kappa}}\right)
$$

Let us denote such unique large cluster as $\mathfrak{C}^{*}$. Of course,

$$
\mathbb{P}_{N}^{\beta, \lambda}\left(x \in \mathfrak{C}^{*}\right)=\mathbb{P}_{N}^{\beta, \lambda}\left(\mathfrak{C}(x)=\mathfrak{C}^{*}\right)=\mathbb{P}_{N}^{\beta, \lambda}(\mathfrak{C}(x) \text { is large })\left(1-O\left(\frac{1}{N^{\kappa}}\right)\right)
$$

does not depend on $x \in \mathfrak{G}_{N}^{\beta}$. In particular, up to $O\left(\frac{1}{N^{\kappa}}\right)$-quantities,

$$
\begin{align*}
\mathbb{E}_{N}^{\beta, \lambda}(\mathcal{M}) & =\mathbb{E}_{N}^{\beta, \lambda}\left(\sum_{i=1}^{N} \int_{0}^{\beta} \mathbb{I}_{\left\{(i, t) \in \mathfrak{C}^{*}\right\}} \mathrm{d} t\right)=\mathbb{E}_{N}^{\beta, \lambda}\left(\sum_{i=1}^{N} \int_{0}^{\beta} \mathbb{I}_{\{\mathfrak{C}((i, t)) \text { is large }\}} \mathrm{d} t\right)  \tag{3.43}\\
& =\sum_{i=1}^{N} \int_{0}^{\beta} \mathbb{P}_{N}^{\beta, \lambda}(\mathfrak{C}((i, t)) \text { is large }) \mathrm{d} t=N \beta p(\beta, \lambda)(1+o(1)),
\end{align*}
$$

where we have used (3.42) in the last step. Accordingly, define $\rho(\beta, \lambda)=N \beta p(\beta, \lambda)$. Let us compute,

$$
\mathbb{E}_{N}^{\beta, \lambda}\left(\mathcal{M}^{2}\right)=\sum_{i=1}^{N} \sum_{j=1}^{N} \int_{0}^{\beta} \int_{0}^{\beta} \mathbb{P}_{N}^{\beta, \lambda}\left((i, t) \in \mathfrak{C}^{*} ;(j, s) \in \mathfrak{C}^{*}\right) \mathrm{d} t \mathrm{~d} s
$$

However, for $i \neq j$ and any $t, s \in[0, \beta)$,

$$
\mathbb{P}_{N}^{\beta, \lambda}\left((i, t) \in \mathfrak{C}^{*} ;(j, s) \in \mathfrak{C}^{*}\right)=\mathbb{P}_{N}^{\beta, \lambda}((i, t) \longleftrightarrow(j, s))(1+o(1))=p(\beta, \lambda)^{2}(1+o(1))
$$

as it follows from the proof of Theorem A in Subsection 3.3 below. Therefore,

$$
\mathbb{E}_{N}^{\beta, \lambda}\left(\mathcal{M}^{2}\right) \leqslant \beta^{2} N+N^{2} \beta^{2} p(\beta, \lambda)^{2}(1+o(1))
$$

As a result,

$$
\operatorname{Var}_{N}^{\beta, \lambda}\left(\frac{\mathcal{M}}{N}\right)=o(1)
$$

and both (1.13) and (1.15) follow.
3.3. LRO: Proof of Theorem A. Define $x=(i, t)$ and $y=(j, s)$, then for $i \neq j$

$$
\mathbb{P}_{N}^{\beta, \lambda}((i, t) \longleftrightarrow(j, s))=\mathbb{P}_{N}^{\beta, \lambda}(\mathfrak{C}(x) \cap \mathfrak{C}(y) \neq \emptyset)
$$

By (3.36),

$$
\mathbb{P}_{N}^{\beta, \lambda}(\mathfrak{C}(x) \cap \mathfrak{C}(y) \neq \emptyset \mid \mathfrak{C}(y) \text { is small }) \leqslant \mathbb{P}_{N}^{\beta, \lambda}\left(j \in \mathcal{S}^{x} \mid \# \mathcal{S}^{x} \leqslant c \log N\right)=O\left(\frac{\log N}{N}\right)
$$

Hence,

$$
\begin{align*}
& \mathbb{P}_{N}^{\beta, \lambda}(\mathfrak{C}(x) \cap \mathfrak{C}(y) \neq \emptyset ; \text { both } \mathfrak{C}(\mathrm{x}) \text { and } \mathfrak{C}(\mathrm{y}) \text { are small }) \\
& \leqslant \mathbb{P}_{N}^{\beta, \lambda}(\mathfrak{C}(x) \cap \mathfrak{C}(y) \neq \emptyset \mid \mathfrak{C}(\mathrm{x}) \text { is small })=O\left(\frac{\log N}{N}\right) \tag{3.44}
\end{align*}
$$

Note that if $\mathfrak{C}(x) \cap \mathfrak{C}(y) \neq \emptyset$, then, up to a $o\left(\frac{1}{N^{\kappa}}\right)$-probability, either both $\mathfrak{C}(x)$ and $\mathfrak{C}(y)$ are large or both $\mathfrak{C}(x)$ and $\mathfrak{C}(y)$ are small. Thus, in view of (3.44),

$$
\begin{aligned}
& \mathbb{P}_{N}^{\beta, \lambda}(\mathfrak{C}(x) \cap \mathfrak{C}(y) \neq \emptyset) \\
& \quad=\mathbb{P}_{N}^{\beta, \lambda}(\mathfrak{C}(x) \cap \mathfrak{C}(y) \neq \emptyset ; \text { both } \mathfrak{C}(\mathrm{x}) \text { and } \mathfrak{C}(\mathrm{y}) \text { are large })+O\left(\frac{\log N}{N}\right) .
\end{aligned}
$$

On the other hand we already know from Lemma 2.2 that two large clusters merge with high $\left(1-o\left(\frac{1}{N^{\kappa}}\right)\right)$ probability. Therefore,

$$
\mathbb{P}_{N}^{\beta, \lambda}(\mathfrak{C}(x) \cap \mathfrak{C}(y) \neq \emptyset)=\mathbb{P}_{N}^{\beta, \lambda}(\text { both } \mathfrak{C}(\mathrm{x}) \text { and } \mathfrak{C}(\mathrm{y}) \text { are large })+O\left(\frac{\log N}{N}\right)
$$

The asymptotic formula (3.42) implies that up to $o(1)$-terms, $\mathbb{P}_{N}^{\beta, \lambda}($ both $\mathfrak{C}(\mathrm{x})$ and $\mathfrak{C}(\mathrm{y})$ are large $)=2 p(\beta, \lambda)-1+\mathbb{P}_{N}^{\beta, \lambda}($ both $\mathfrak{C}(\mathrm{x})$ and $\mathfrak{C}(\mathrm{y})$ are small $)$. Since both $\{\mathfrak{C}(\mathrm{x})$ is small $\}$ and $\{\mathfrak{C}(\mathrm{x})$ is small $\}$ are decreasing events, by the FKG properties of $\mathbb{P}_{N}^{\beta, \lambda}(\cdot)$,

$$
\mathbb{P}_{N}^{\beta, \lambda}(\text { both } \mathfrak{C}(\mathrm{x}) \text { and } \mathfrak{C}(\mathrm{y}) \text { are small }) \geqslant(1-p(\beta, \lambda))^{2}(1+o(1))
$$

To get a complimentary upper bound, notice that

$$
\begin{aligned}
& \mathbb{P}_{N}^{\beta, \lambda}\left(\mathcal{S}^{x} \cap \mathcal{S}^{y} \neq \emptyset \mid \text { both } \mathfrak{C}(\mathrm{x}) \text { and } \mathfrak{C}(\mathrm{y}) \text { are small }\right) \\
& \leqslant \mathbb{P}_{N}^{\beta, \lambda}\left(\mathcal{S}^{x} \cap \mathcal{S}^{y} \backslash\{j\} \neq \emptyset \mid \# \mathcal{S}^{x} \leqslant c \log N ; \# \mathcal{S}^{y} \leqslant c \log N\right) \\
& \leqslant 2 c \log N \mathbb{P}_{N}^{\beta, \lambda}\left(i \in \mathcal{S}^{y} \mid \# \mathcal{S}^{y} \leqslant c \log N\right)=O\left(\frac{(\log N)^{2}}{N}\right)
\end{aligned}
$$

as it follows from (3.36) and $\mathrm{G}_{N}$-permutation invariance. Thus, $\mathbb{P}_{N}^{\beta, \lambda}($ both $\mathfrak{C}(\mathrm{x})$ and $\mathfrak{C}(\mathrm{y})$ are small $)$
$=\mathbb{P}_{N}^{\beta, \lambda}\left(\right.$ both $\mathfrak{C}(\mathrm{x})$ and $\mathfrak{C}(\mathrm{y})$ are small; $\left.\mathcal{S}^{x} \cap \mathcal{S}^{y}=\emptyset\right)+O\left(\frac{(\log N)^{2}}{N}\right)$
$\leqslant \mathbb{P}_{N}^{\beta, \lambda}\left(\mathfrak{C}(\mathrm{x})\right.$ is small $\mid \mathcal{S}^{x} \cap \mathcal{S}^{y}=\emptyset ; \mathfrak{C}(\mathrm{y})$ is small $) \mathbb{P}_{N}^{\beta, \lambda}(\mathfrak{C}(\mathrm{y})$ is small $)+O\left(\frac{(\log N)^{2}}{N}\right)$

However,

$$
\mathbb{P}_{N}^{\beta, \lambda}\left(\mathfrak{C}(\mathrm{x}) \text { is small } \mid \mathcal{S}^{x} \cap \mathcal{S}^{y}=\emptyset ; \mathfrak{C}(\mathrm{y}) \text { is small }\right) \leqslant \mathbb{P}_{N-c \log N}^{\tilde{\beta}, \lambda}(\mathfrak{C}(\mathrm{x}) \text { is small })
$$

where $\tilde{\beta}=\beta \frac{N}{N-c \log N}$.
As $N$ tends to $\infty$, the right hand side above converges to $1-p(\beta, \lambda)$. Thereby, we are able to conclude,

$$
\mathbb{P}_{N}^{\beta, \lambda}(\text { both } \mathfrak{C}(\mathrm{x}) \text { and } \mathfrak{C}(\mathrm{y}) \text { are small })=(1-p(\beta, \lambda))^{2}(1+o(1))
$$

and (1.8) follows.
Acknowledgments D.I. thanks Petra Scudo, for an information about quantum Ising type models, and Errico Presutti for several useful discussions and a suggestion to study mean field models first.

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[^0]:    ${ }^{0}$ AMS 2000 Subject Classification: 60 J 80 , $60 \mathrm{~K} 45,82 \mathrm{~B} 10,82 \mathrm{~B} 20,82 \mathrm{~B} 26$
    Key Words and Phrases: Quantum Curie-Weiss model, FK representation, percolation, giant components of random graphs, branching random walks

    Date: March 15, 2006.
    Partly supported by the EU Network Postdoctoral Training Program in Mathematical Analysis of Large Quantum Systems under the contract HPRN-CT-2002-00277.

